

# LOVÁSZ LOCAL LEMMA

Events  $A_1, A_2, \dots, A_n \rightarrow$  Can these be avoided?

Case 1:  $A_i$ 's mutually independent and  $P(A_i) < 1 \forall i \rightarrow$

$$P(\bigwedge_i \bar{A}_i) = \prod_i (1 - P(A_i)) > 0 \checkmark$$

Union Bound  $\rightarrow$  Case 2:  $A_i$ 's not necessarily mutually independent but having very small probability  $\rightarrow$

$$P(\bigwedge_i \bar{A}_i) = 1 - P(\bigvee_i A_i) > 1 - \sum_i P(A_i) > 0 \checkmark$$

But, are cases always so nice?

We might be given events with moderately small probabilities, and a certain level of dependence.

So is there a chance of avoiding the events?

- Answered by Lovász - Erdős (1975) & improved by Lovász (1977).

★ Dependency Graph:

An undirected graph  $G$

$$\exists V_G = \{A_1, A_2, \dots, A_n\}$$

and  $A_i$  and  $A_j$  are

dependent  $\Rightarrow \{A_i, A_j\} \in E_G$ .



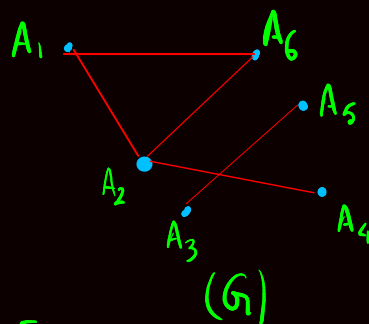
Mind it - the converse is not true

and the dependency graph is not unique.

For example, if all  $A_i$ 's are independent,

$E_G = \emptyset$  is valid, whereas a complete  $K_n$  is

valid in any case but our job is to simplify it as much as possible.



(define degree for first yrs)

(1) Lemma 1: Lovász-Erdős (1975) Symmetric version:

Conditions: (1) Each event dependent on at most  $d > 0$  other events ( $\Delta_G \leq d$ ). → Hence Local

Sufficient only (2)  $P[A_i] \leq \frac{1}{4d} \forall i \Rightarrow P[\bigwedge_j \bar{A}_j] > 0$

Here we have a limit  $d$ . Notice that when  $d = n-1$  (all events are dependent on each other), we get back our union bound.

$$\left( n \cdot \frac{1}{4(n-1)} < 1 \Leftrightarrow n \geq 2 \right)$$

In order to prove this, we make a claim. You choose an event  $A_i$  and take any subset of events  $T$ . The probability of  $A_i$  happening given that none of the events in  $T$  occurred is at most twice the unconditional probability of  $A_i$  occurring.

Mathematically,  $P[A_i | \bigwedge_{j \in T} \bar{A}_j] \leq 2P[A_i] \forall T \subseteq V_G$ .

We prove it using induction.

Base case:  $|T| = 1$ .

$$P[A_1 | \bar{A}_2] = \frac{P[A_1 \cap \bar{A}_2]}{1 - P[A_2]} \leq \frac{P[A_1]}{1 - P[A_2]} \leq 2P[A_1] \left( \because P[A_2] \leq \frac{1}{4d} < \frac{1}{2} \right)$$

Inductive case: Let the claim be true  $\forall |T| \leq k$ .

For  $|T| = k+1$ , we divide  $T \rightarrow T_1$  (joined to  $A_i$  in  $G$ ) - nearby events ( $\leq d$ )  
 $\rightarrow T_2$  (others) - far away events

$$P[A_i | \bigwedge_j \bar{A}_j] = \frac{P[A_i \cap \bigwedge_{j \in T} \bar{A}_j]}{P[\bigwedge_{j \in T} \bar{A}_j]} = \frac{P[A_i \cap \bigwedge_{j \in T_1} \bar{A}_j \cap \bigwedge_{j \in T_2} \bar{A}_j]}{P[\bigwedge_{j \in T_1} \bar{A}_j \cap \bigwedge_{j \in T_2} \bar{A}_j]}$$

$$= \left( \frac{P[A_i \cap \bigwedge_{j \in T_1} \bar{A}_j \cap \bigwedge_{j \in T_2} \bar{A}_j]}{P[\bigwedge_{j \in T_2} \bar{A}_j]} \right) / \left( \frac{P[\bigwedge_{j \in T_1} \bar{A}_j \cap \bigwedge_{j \in T_2} \bar{A}_j]}{P[\bigwedge_{j \in T_2} \bar{A}_j]} \right) = \frac{P[A_i \cap \bigwedge_{j \in T_1} \bar{A}_j | \bigwedge_{j \in T_2} \bar{A}_j]}{P[\bigwedge_{j \in T_1} \bar{A}_j | \bigwedge_{j \in T_2} \bar{A}_j]}$$

$$\leq \frac{P[A_i]}{1 - P[\bigvee_{j \in T_1} A_j | \bigwedge_{j \in T_2} \bar{A}_j]} \leq \frac{P[A_i]}{1 - \sum_{j \in T_1} P[A_j | \bigwedge_{j \in T_2} \bar{A}_j]} \leq 2P[A_i]. \left( \because \sum_{j \in T_1} P[A_j | \bigwedge_{j \in T_2} \bar{A}_j] \leq 2 \cdot d \cdot \frac{1}{4d} = \frac{1}{2} \right)$$

$|T_1| \leq d \leq 2P[A_j]$  by inductive claim.

Hence, claim proved by induction.

## Proof of Erdős-Lovász:

$$\begin{aligned}
 P\left[\bigwedge_{i=1}^n \bar{A}_i\right] &= P[\bar{A}_1] \cdot P[\bar{A}_2 | \bar{A}_1] \cdot P[\bar{A}_3 | \bar{A}_1 \wedge \bar{A}_2] \cdot \dots \cdot P[\bar{A}_n | \bigwedge_{i=1}^{n-1} \bar{A}_i] \\
 &= (1 - P[A_1]) \cdot (1 - P[A_2 | \bar{A}_1]) \cdot (1 - P[A_3 | \bar{A}_1 \wedge \bar{A}_2]) \cdot \dots \cdot (1 - P[A_n | \bigwedge_{i=1}^{n-1} \bar{A}_i]) \\
 &\geq (1 - 2P[A_1]) \cdot (1 - 2P[A_2]) \cdot \dots \cdot (1 - 2P[A_n]) > 0 \quad (\because P[A_i] < \frac{1}{2} \forall i).
 \end{aligned}$$

## Lemma 2: Asymmetric Version of Lemma 1:

If  $\exists \chi_1, \chi_2, \dots, \chi_n \in \mathbb{R}[0, 1) \ni P[A_i] \leq \chi_i \cdot \prod_{\{A_j, A_k\} \in E_G} (1 - \chi_j) \forall i$ , then:

$$P\left(\bigwedge \bar{A}_i\right) \geq \prod_i (1 - \chi_i).$$

$\{A_i, A_j\} \in E_G$   
↪ nearby events  $T_i$

Proof: Similar to proof of lemma 1.

Claim:  $P[A_i | \bigwedge_T \bar{A}_j] \leq \chi_i \quad \forall T \subseteq V_G$ .

Let  $T_1 = \{B_1, B_2, \dots, B_k\}$ .

By Bayes,

$$P\left(\bigwedge_{T_1} \bar{A}_j \mid \bigwedge_{T_2} \bar{A}_j\right) = P(\bar{B}_1 \mid \bigwedge_{t=2}^k \bar{B}_t \wedge \bigwedge_{T_2} \bar{A}_j) \cdot P(\bar{B}_2 \mid \bigwedge_{t=3}^k \bar{B}_t \wedge \bigwedge_{T_2} \bar{A}_j) \cdot \dots \cdot P(\bar{B}_k \mid \bigwedge_{T_2} \bar{A}_j)$$

Inductive assumption  $\geq \prod_{T_1} (1 - \chi_j)$

$$P[A_1 | \bigwedge_T \bar{A}_j] = \frac{P[A_1 \cap \bigwedge_{T_1} \bar{A}_j \mid \bigwedge_{T_2} \bar{A}_j]}{P[\bigwedge_{T_1} \bar{A}_j \mid \bigwedge_{T_2} \bar{A}_j]} \leq \frac{P[A_1]}{P[\bigwedge_{T_1} \bar{A}_j \mid \bigwedge_{T_2} \bar{A}_j]} \leq \frac{\chi_1 \cdot \prod_{T_1} (1 - \chi_j)}{\prod_{T_1} (1 - \chi_j)} = \chi_1.$$

Hence  $P(\bigwedge \bar{A}_j) \geq \prod_i (1 - \chi_i)$  (Expand like we did in lemma 1).

## ★ Lemma 3: Lovász Local Lemma (1977)

Conditions: (1) Each event dependent on at most  $d$  other events

Sufficient  $\leftarrow$  (2)  $P[A_i] \leq \frac{1}{e(d+1)} \quad \forall i$ . (More relaxed than Lemma 1 for  $d > 2$ ).

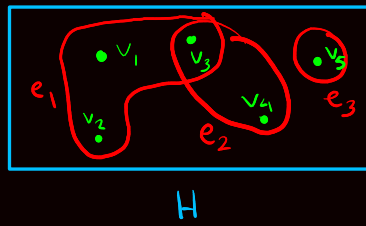
Proof: Given condition 1 is followed,  $\chi_i \prod_{T_i} (1 - \chi_j)$  attains its maximum when  $\chi_i = \frac{1}{d+1} \quad \forall i$  (Check using symmetry & calculus).

This condition translates to:

$$\begin{aligned}
 \chi_i \prod_{T_i} (1 - \chi_j) &= \frac{1}{d+1} \left(\frac{d}{d+1}\right)^{|T_i|} \geq \frac{1}{d+1} \left(\frac{d}{d+1}\right)^d \geq \frac{1}{e(d+1)} \geq P[A_i]. \quad \left(\because \left(\frac{d}{d+1}\right)^d \downarrow \frac{1}{e} \text{ as } d \rightarrow \infty\right). \\
 &(\because |T_i| \leq d \text{ \& decreasing function}) \Rightarrow \text{Conditions for Lemma 2 are met.}
 \end{aligned}$$

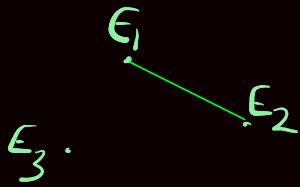
# What is a hypergraph?

It's just a generalized graph where an edge is a subset of any number of vertices!



See how it might relate to a dependency graph?

Edge set of hypergraph →  
 Vertex set of dependency graph  
 Join vertices if edge sets intersect!



Now, let's go the other way round! (Somewhat)

You have a dependency graph with  $n$  vertices  $A_1, \dots, A_n$  (events).

Say each event has independent subevents  $(A_{11}, A_{12}, \dots, A_{1n_1})$ ,

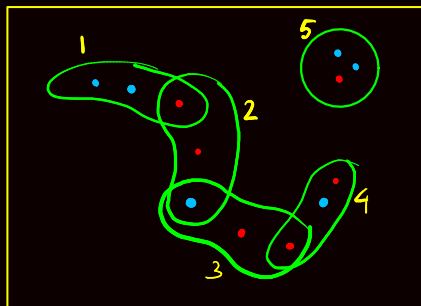
$(A_{21}, A_{22}, \dots, A_{2n_2}), \dots$  and so on.

We split the vertices into the independent subevents and "encapsulate" them into edge sets. Notice that two events with  $\rightarrow \geq 2$  per event have a common subevent  $\Leftrightarrow$  they are dependent  $\Leftrightarrow$  the edge sets intersect  $\Rightarrow$  the events are joined by an edge in the dependency graph. This gives us our hypergraph!

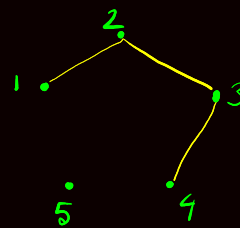
## Applications:

Let in a hypergraph with  $n$  edge sets,  $A_i$  be the event that the  $i^{\text{th}}$  edge set is monochromatic (bad event we want to avoid).

R & B



3-uniform hypergraph



Dependency graph

$P(A_i) = ?$

Conditions met?

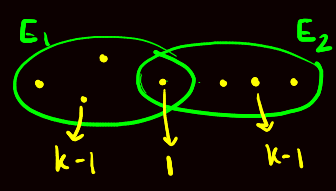
→ 2-Colouring of k-uniform Hypergraph: Note:  $k \geq 2$ .

↘ Each edge set has k vertices

Q1: Given an edge set with k vertices, what's the probability that it's monochromatic?

Choose what to colour the vertices with in 2 ways (R or B).  
 Then all the vertices will be coloured in that in  $(\frac{1}{2})^k$  ways.  
 $\Rightarrow P(A_i) = 2^{-(k-1)}$ .

Q2: Say  $E_1 \times E_2$  two edge sets with one intersecting vertex. Given one edge set is monochromatic, what is the probability that the other edge set is also monochromatic?



Say  $E_1$  monochromatic  $\Rightarrow$  the intersecting vertex fixes the colour of the remaining  $k-1$  vertices of  $E_2$   
 $\Rightarrow 2^{-(k-1)}$ .

$\therefore P(A_1 | A_2) = P(A_1) \Rightarrow A_1 \times A_2$  are independent events.

Q3: What is the minimum number of edges  $m(k)$  required for a k-uniform hypergraph to not be two-colourable?

Say  $k=2$  (just a graph).

- 1 edge:
- 2 edges:
- 3 edges:

$\therefore m(2) = 3$

A closed form of  $m(k)$  is still open, although  $m(k) \geq 2^{k-1}$ .

Claim:  $m(k) \geq 2^{k-1}$

Take any k-uniform hypergraph H with  $m < 2^{k-1}$  edges.

Assign vertex colours randomly.

$P(e \text{ is monochromatic}) = 2^{-(k-1)}$

$P(\text{an edge is monochromatic}) = P(\bigvee_i A_i) \leq \sum_i P(A_i) = \frac{m}{2^{k-1}} < 1$ .

$\therefore P(\text{no edge is monochromatic}) = P(\overline{\bigvee_i A_i}) > 0$ .

$\therefore H$  is two-colourable.  $\therefore$  All hypergraphs with  $m < 2^{k-1}$  edges are two-colourable

What does the Lovász Lemma tell us? (Page 4 - show not necessary)

Thm: Let  $H$  be a  $k$ -uniform hypergraph  $\ni$  each edge intersects almost  $d$  others. Then  $d \leq \frac{2^{k-1}}{e} - 1 \Rightarrow H$  is two-colourable.

Pf: Do it yourself (Hint:  $\mathbb{P}[A_i] = 2^{-(k-1)}$ ).

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Reference:  $\rightarrow$  Talk on Lovász Local Lemma by

Profesor Jaikumar Radhakrishnan.

$\rightarrow$  Lecture notes of Math 233A by Stanford University

(Will be shared in group).

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